

Fall 2010 Graduate Course on
Dynamic Learning

Chapter 4: Particle Filters

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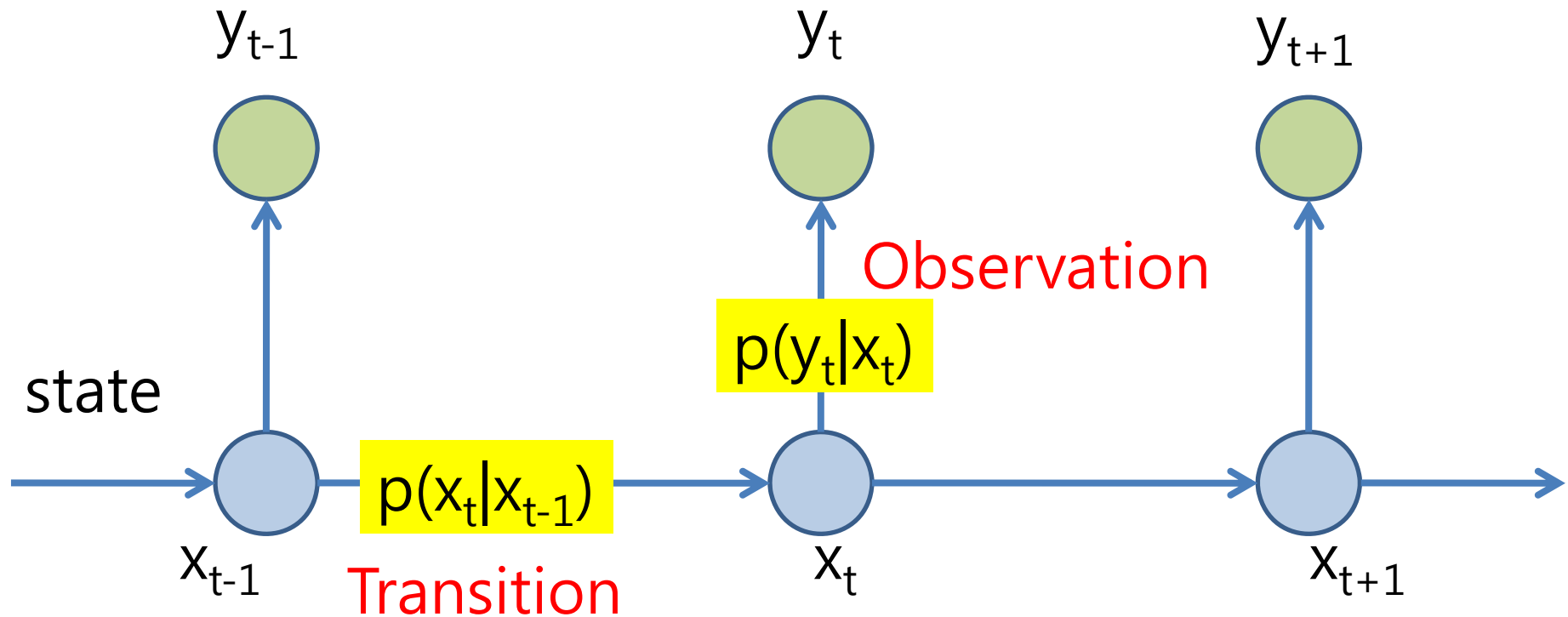
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Overview

- Filtering Problem
- Sequential Bayesian Filtering
- Particle Filter
- Monte Carlo (MC) Approximation
- MC with Importance Sampling (IS)
- Sequential Importance Sampling (SIS)
- Sampling Importance Resampling (SIR)

Filtering / Tracking

- We want to track the unknown state x of a system as it evolves over time based on the (noisy) observations y that arrive sequentially.



Dynamical System

x_t is state vector at time t , y_t is observations at time t

State equation $p(x_t | x_{t-1})$

Observation equation $p(y_t | x_t)$

Note: The forms of $p(x_t | x_{t-1})$ and $p(y_t | x_t)$ depend on the state transition function $f_X(\cdot)$ and observation function $f_Y(\cdot)$.

State equation: $x_t = f_X(x_{t-1}, u_t)$

f_X state transition function

u_t process noise with known distribution

Observation equation: $y_t = f_Y(x_t, v_t)$

f_Y observation function

v_t observation noise with known distribution

Filtering Problem

- The objective is to estimate unknown state x_t , based on a sequence of observations y_t , $t=0,1, \dots$

Find posterior distribution $p(x_{0:t} | y_{1:t})$

- By knowing posterior distribution (of the states) a number of estimates can be computed, e.g. the expected value of some function $f(\cdot)$ that depends on the state values:

$$E[f(x_{0:t})] = \int f(x_{0:t}) p(x_{0:t} | y_{1:t}) dx_{0:t}$$

Formally...

Let:

State vector $x_{0:t} = (x_0, \dots, x_t)$

Observation vector $y_{1:t} = (y_1, \dots, y_t)$

Find:

PDF $p(x_{0:t} | y_{1:t})$ posterior distribution

or $p(x_t | y_{1:t})$ filtering distribution

Given:

$p(x_0)$ prior distribution (on state)

$p(x_t | x_{t-1})$ transition probability (e.g., motor model)

$p(y_t | x_t)$ observation probability (e.g., sensor model)

$p(x_0)$ is given.

$t = 0$, observe y_0 .

$$\text{Update} \quad p(x_0 | y_0) = \frac{p(y_0 | x_0)}{p(y_0)} p(x_0) \quad (\text{Bayes theorem})$$

$$\text{Predict} \quad p(x_1 | y_0) = \int p(x_1 | x_0) p(x_0 | y_0) dx_0 \quad (\text{Markovian})$$

$t = 1$, observe y_1 from x_1

$$\text{Update} \quad p(x_1 | y_1) = \frac{p(y_1 | x_1)}{p(y_1)} p(x_1)$$

$$\text{Predict} \quad p(x_2 | y_1) = \int p(x_2 | x_1) p(x_1 | y_1) dx_1$$

$t = 2$, observe y_2 from x_2

$$\text{Update} \quad p(x_2 | y_{1:2}) = \frac{p(y_2 | x_2)}{p(y_2 | y_1)} p(x_2 | y_1)$$

$$\text{Predict} \quad p(x_3 | y_{1:2}) = \int p(x_3 | x_2) p(x_2 | y_{1:2}) dx_2$$

$t = 3$, observe y_3 from x_3

$$\text{Update} \quad p(x_3 | y_{1:3}) = \frac{p(y_3 | x_3)}{p(y_3 | y_{1:2})} p(x_3 | y_{1:2})$$

$$\text{Predict} \quad p(x_4 | y_{1:3}) = \int p(x_4 | x_3) p(x_3 | y_{1:3}) dx_3$$

Sequential Bayesian Filtering

Given $p(x_{t-1} | y_{1:t-1})$... prior (filtering) distribution (i.e., before observing y_t)

1. Prediction

$$p(x_t | y_{1:t-1}) = \int_{x_{t-1}} p(x_t | x_{t-1}) p(x_{t-1} | y_{1:t-1}) dx_{t-1} \quad (\text{Eqn. 1})$$

$$\begin{aligned} \text{since } p(x_t | y_{1:t-1}) &= \int_{x_{t-1}} p(x_t, x_{t-1} | y_{1:t-1}) dx_{t-1} \\ &= \int_{x_{t-1}} p(x_t | x_{t-1}, y_{1:t-1}) p(x_{t-1} | y_{1:t-1}) dx_{t-1} \end{aligned}$$

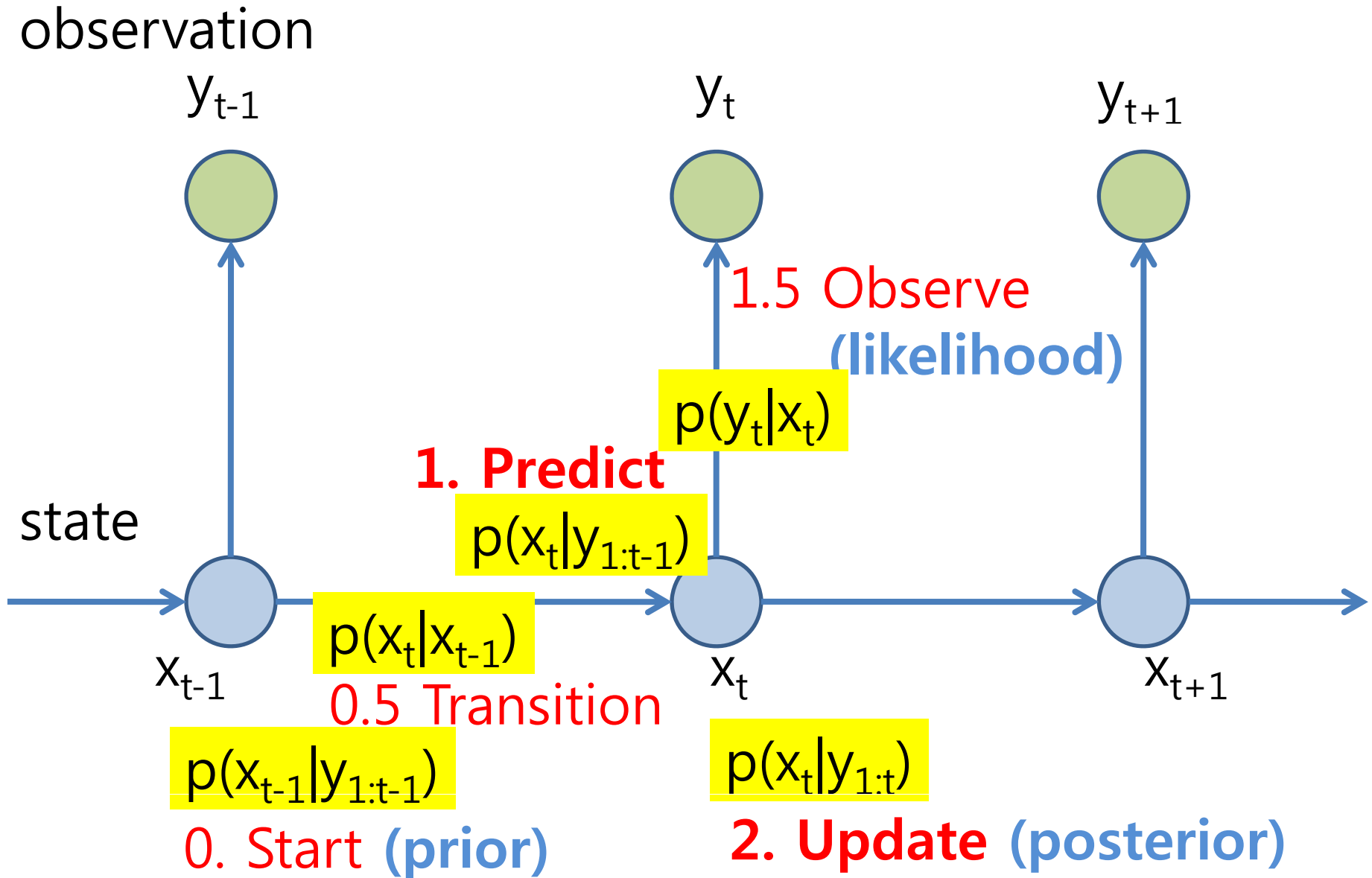
$$\text{note: } p(a) = \int_b p(a, b) db \quad \text{and} \quad p(a, b | c) = p(a | b, c) p(b | c)$$

2. Update ... posterior distribution (after observing y_t)

$$p(x_t | y_{1:t}) = \frac{p(y_t | x_t) p(x_t | y_{1:t-1})}{p(y_t | y_{1:t-1})} \quad (\text{Eqn. 2})$$

$$\text{where } p(y_t | y_{1:t-1}) = \int_{x_t} p(y_t | x_t) p(x_t | y_{1:t-1}) dx_t$$

Graphically...



A Special Case: Kalman Filter

$$p(x_{t-1} | y_{1:t-1}) = N(x_{t-1} | m_{t-1|t-1}, P_{t-1|t-1})$$

$$p(x_t | y_{1:t-1}) = N(x_t | m_{t|t-1}, P_{t|t-1})$$

$$p(x_t | y_{1:t}) = N(x_t | m_{t|t}, P_{t|t})$$

$$m_{t|t-1} = F_t m_{t-1|t-1}$$

$$P_{t|t-1} = Q_{t-1} + F_t P_{t-1|t-1} F_t^T$$

...

$$x_t = F_t x_{t-1} + v_{t-1}, \quad v_{t-1} \sim N(0, Q_{t-1}) \quad \dots \text{linear and Gaussian}$$

$$y_t = H_t x_t + n_t, \quad n_t \sim N(0, R_t)$$

F_t : transition matrix (known)

H_t : observation matrix (known)

Particle Filters

- Particle filter is a technique for implementing recursive Bayesian filter by Monte Carlo sampling
- The idea is to represent the posterior density by a set of random samples (particles) with associated weights.
 - Compute estimates based on these samples and weights.
- Many different names....
 - Sequential Monte Carlo (SMC)
 - Condensation method
 - Survival of the fittest (evolutionary computation?)

Advantages of Particle Filters

- Ability to represent arbitrary densities
 - Can deal with non-linearities
 - Non-Gaussian noise
- Particle filters focus adaptively on probable regions of state space
 - In contrast, HMM filters discretize the state space to N fixed states.
- Can be implemented in $O(Ns)$
 - Ns : sample size
 - Easy to implement
 - Easy to parallelize

Sample-Based PDF Representation

- Monte Carlo characterization of pdf
- Represent posterior density by a set of random i.i.d. samples (particles) from the pdf $p(x_{0:t}|y_{1:t})$
- For large number N of particles equivalent to functional description of pdf
- For $N \rightarrow \infty$, Monte Carlo method approaches optimal Bayesian estimate.

Monte Carlo (MC) Approximation

$$E_p[f(x)] = \int p(x) f(x) dx$$

$$\approx \frac{1}{N} \sum_{i=1}^N f(x^{(i)}), \quad x^{(i)} \sim p(x) = N(0, \sigma^2)$$

- Monte Carlo approach
 1. Simulate N random variables from $p(x)$, e.g. Normal distribution

$$x^{(i)} \sim p(x) = N(0, \sigma^2)$$

2. Compute the average

$$E_p[f(x)] = \frac{1}{N} \sum_{i=1}^N f(x^{(i)}),$$

MC with Importance Sampling

$$\begin{aligned} E_p[f(x)] &= \int_x p(x) f(x) dx \\ &= \int_x \frac{p(x)}{q(x)} q(x) f(x) dx \\ &\approx \sum_{i=1}^N w_i f(x^{(i)}) \end{aligned}$$

$x^{(i)} \sim q(x)$ $q(x)$: proposal distribution

$w_i = \frac{p(x^{(i)})}{q(x^{(i)})}$ w_i : importance weight

Note: $q(x)$ is easier to sample from than $p(x)$.

Importance Sampling (IS)

$$E[f(x_{0:t})] = \int f(x_{0:t}) p(x_{0:t} | y_{1:t}) dx_{0:t}$$

$$\approx \sum_{i=1}^N w_i f(x_{0:t}^{(i)})$$

$$x_{0:t}^{(i)} \sim q(x_{0:t} | y_{1:t}) \quad q(x): \text{proposal distribution}$$

$$w_i = \frac{p(x_{0:t}^{(i)} | y_{1:t})}{q(x_{0:t}^{(i)} | y_{1:t})} \quad w_i: \text{importance weight}$$

Importance Sampling: Procedure

1. Draw N samples $x_{0:t}^{(i)}$ from proposal distribution $q(\cdot)$.

$$x_{0:t}^{(i)} \sim q(x_{0:t} | y_{1:t})$$

2. Compute importance weight

$$w(x_{0:t}^{(i)}) = \frac{p(x_{0:t}^{(i)} | y_{1:t})}{q(x_{0:t}^{(i)} | y_{1:t})}$$

3. Estimate an arbitrary function $f(\cdot)$:

$$E[f(x_{0:t} | y_{1:t})] \approx \sum_{i=1}^N f(x_{0:t}^{(i)}) \tilde{w}_t^{(i)}, \quad \tilde{w}_t^{(i)} = \frac{w(x_{0:t}^{(i)})}{\sum_{j=1}^N w(x_{0:t}^{(j)})}$$

Sequential Importance Sampling (SIS): Recursive Estimation

Augmenting the samples

$$\begin{aligned}q(x_{0:t} | y_{1:t}) &= q(x_{0:t-1} | y_{1:t-1})q(x_t | x_{0:t-1}, y_{1:t}) \\ &= q(x_{0:t-1} | y_{1:t-1})q(x_t | x_{t-1}, y_t)\end{aligned}$$

$$x_t^{(i)} \sim q(x_t | x_{t-1}, y_t)$$

$$\text{(cf. non-sequential IS: } x_t^{(i)} \sim q(x_{0:t} | y_{1:t}))$$

Weight update

$$w_t^{(i)} \propto w_{t-1}^{(i)} \frac{p(y_t | x_t^{(i)})p(x_t^{(i)} | x_{t-1}^{(i)})}{q(x_t^{(i)} | x_{t-1}^{(i)}, y_t)}$$

Sequential Importance Sampling: Idea

- Update filtering density using Bayesian filtering
- Compute integrals using importance sampling
- The **filtering density** $p(x_t | y_{1:t})$ is represented using particles and their weights

$$\{x_t^{(i)}, w_t^{(i)}\}_{i=1}^N$$

- Compute weights using:

$$w_t^{(i)} = \frac{p(x_t^{(i)}, y_{1:t})}{q(x_t^{(i)}, y_{1:t})}$$

Sequential Importance Sampling: Procedure

1. Particle generation $x_t^{(i)} \sim q(x_t | x_{t-1}^{(i)}, y_t) = p(x_t | x_{t-1}^{(i)})$

2. Weight computation $w_t^{(i)} = w_{t-1}^{(i)} p(y_t | x_t^{(i)})$

Weight normalization $\tilde{w}_t^{(i)} = \frac{w_t^{(i)}}{\sum_{j=1}^N w_t^{(j)}}$

3. Estimation computation $E[f(x_t | y_{1:t})] = \sum_{i=1}^N f(x_t^{(i)}) \tilde{w}_t^{(i)}$

Note: Step 1 above assumes the proposal density to be the prior.

This does not use the information from observations. Alternatively, the proposal density could be

$$x_t^{(i)} \sim q(x_t | x_{t-1}^{(i)}, y_t) = p(x_t | x_{t-1}^{(i)}, y_t)$$

that minimizes the variance of w_t (Doucet et al., 1999).

Resampling

- SIS suffers from degeneracy problems, i.e. a small number of particles have big weights and the rest have extremely small values.
- Remedy: SIR introduces a selection (resampling) step to eliminate samples with low importance ratios (weights) and multiply samples with high importance ratios.
- Resampling maps the weighted random measure on to the equally weighted random measure by sampling uniformly with replacement from $\{x_{0:t}^{(i)}\}_{i=1}^N$ with probabilities $\{w_t^{(i)}\}_{i=1}^N$:

$$\{\tilde{x}_{0:t}^{(i)}, N^{-1}\}_{i=1}^N \sim \{x_{0:t}^{(i)}, w_t^{(i)}(x_{0:t}^{(i)})\}_{i=1}^N$$

Sampling Importance Resampling (SIR) = Sequential Monte Carlo = Particle Filter

1. Initialize $t \leftarrow 0$

- For $i = 1, \dots, N$: sample $x_t^{(i)} \sim p(x_0)$, $t \leftarrow 1$.

2. Importance sampling

- For $i = 1, \dots, N$: sample $x_t^{(i)} \sim q(x_t | x_{t-1}^{(i)}, y_t) = p(x_t | x_{t-1}^{(i)})$

Let $x_{0:t}^{(i)} \triangleq (x_{0:t-1}^{(i)}, x_t^{(i)})$

- For $i = 1, \dots, N$: compute weights $w_t^{(i)} = p(y_t | x_t^{(i)})$

- Normalize the weights: $\tilde{w}_t^{(i)} = w_t^{(i)} / \sum_{j=1}^N w_t^{(j)}$

3. Resampling

- Resample with replacement N particles $x_{0:t}^{(i)}$ according to the importance weights $w_t^{(i)}$, resulting in $\{\tilde{x}_{0:t}^{(i)}, N^{-1}\}_{i=1}^N$.

- New particle population $\{x_{0:t}^{(i)}\}_{i=1}^N \leftarrow \{\tilde{x}_{0:t}^{(i)}\}_{i=1}^N$.

- Set $t \leftarrow t + 1$ and go to step 2.

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